## Probability in Computing



## Reminders

- HW 6 is due Thursday Reading
- LLM 18.18-18.19, 19.2

P 1.4.1, 3.1.4

## Lecture 12

Last time

- Bayes' Rule Example
- Review


## Today

- Pairwise and Mutual Independence


## Midterm Announcement

- Midterm will be Wednesday, March 15, 6:30-8:30 in CGS 129
- Closed book, no devices.
- One hand-written 2-sided page of notes allowed on colored paper, significantly different from white. No mechanical reproduction (writing on a computer and then printing is not allowed).
- All material up to tomorrow's discussion (Friday, 3/2)
- Practice exam problems will be distributed soon; Friday discussion will be practice problems, Tuesday lecture (3/14) will be review.
- Review TopHat, homework, and discussion problems, plus exercises suggested from the 2 nd textbook.


## Reminder: Independent Events

## Definition: Independent Events

Two events $A$ and $B$ are independent if

$$
\operatorname{Pr}(A \cap B)=\operatorname{Pr}(A) \cdot \operatorname{Pr}(B)
$$

- Independent $\neq$ disjoint! When events A and B are disjoint: if A happens then B is guaranteed not to happen.


## Top Hat question (Join Code: 033357)

Let $A$ and $B$ be two disjoint events with $\operatorname{Pr}(A)>0$ and $\operatorname{Pr}(B)>0$ Are $A$ and $B$ independent?
A. YES, for all $A, B$ satisfying the conditions.
B. NO, for all $A, B$ satisfying the conditions.
C. YES for some $A, B$ satisfying the conditions and NO for others.

## Independence for Three Events

How should we define independence for three events $A, B, C$ ?

- Options

1) $A$ and $B$ are independent: $\operatorname{Pr}(A \cap B)=\operatorname{Pr}(A) \cdot \operatorname{Pr}(B)$
$B$ and $C$ are independent: $\operatorname{Pr}(B \cap C)=\operatorname{Pr}(B) \cdot \operatorname{Pr}(C)$
$A$ and $C$ are independent: $\operatorname{Pr}(A \cap C)=\operatorname{Pr}(A) \cdot \operatorname{Pr}(C)$
2) $\operatorname{Pr}(A \cap B \cap C)=\operatorname{Pr}(A) \cdot \operatorname{Pr}(B) \cdot \operatorname{Pr}(C)$
3) Require both (1) and (2).

Questions: Does (1) imply (2)? Does (2) imply (1)?

## Independence for Three Events

Option (1) does not imply Option (2)

- Experiment: toss a fair coin twice
- $A=$ event that $1^{\text {st }}$ toss is $H$
- $B=$ event that $2^{\text {nd }}$ toss is $H$
$\qquad$
$=\{\mathrm{HH}, \mathrm{HT}\}$
$=\{\mathrm{HH}, \mathrm{TH}\}$
- $C=$ event that tosses are different $=\{H T, T H\}$


## Independence for Three Events

Option (2) does not imply Option (1)

- Experiment: roll a fair die twice
- $A=$ event that $1^{\text {st }}$ roll is 1,2 or 3
- $B=$ event that $1^{\text {st }}$ roll is 3,4 or 5
- $C=$ the sum of the two rolls is $9=\{(3,6),(6,3),(4,5),(5,4)\}$

$$
\operatorname{Pr}(A \cap B \cap C)=
$$

$\operatorname{Pr}(A \cap B)=$

## Independent Events

## Definition: Independent Events

Three events $A, B$, and $C$ are pairwise-independent if

$$
\begin{aligned}
& \operatorname{Pr}(A \cap B)=\operatorname{Pr}(A) \cdot \operatorname{Pr}(B) ; \\
& \operatorname{Pr}(B \cap C)=\operatorname{Pr}(B) \cdot \operatorname{Pr}(C) ; \\
& \operatorname{Pr}(A \cap C)=\operatorname{Pr}(A) \cdot \operatorname{Pr}(C) .
\end{aligned}
$$

Three events $A, B$ and $C$ are mutually independent if they are pairwise independent and

$$
\operatorname{Pr}(A \cap B \cap C)=\operatorname{Pr}(A) \cdot \operatorname{Pr}(B) \cdot \operatorname{Pr}(C) .
$$

## Compact Notation: Sum and Product

- $\Sigma$ : Sigma, summation notation LaTex: \sum
- П: Pi, product notation

LaTex: \prod

$$
\begin{aligned}
& \sum_{i=2}^{5} i=2+3+4+5 \\
& \prod_{n=1}^{6} x=x^{6} \\
& \prod_{i=1}^{2} \prod_{j=4}^{6} 3 i j=\prod_{i=1}^{n}((3 i \cdot 4)(3 i \cdot 5)(3 i \cdot 6)) \\
& =(3 \cdot 1 \cdot 4)(3 \cdot 1 \cdot 5)(3 \cdot 1 \cdot 6) \cdot(3 \cdot 2 \cdot 4)(3 \cdot 2 \cdot 5)(3 \cdot 2 \cdot 6)
\end{aligned}
$$

## Compact Notation with Sets

- $\Sigma$ : Sigma, summation notation LaTex: \sum
- $\Pi: \mathrm{Pi}$, product notation


## LaTex: \prod

 Let $S=\{1,2,3\}$. Then$\sum_{i \in S} i^{2}=$
$\prod_{i \in S} \operatorname{Pr}\left(A_{i}\right)=$
$\bigcap_{i \in S} A_{i}=$

## Independent Events

## Definition: Independent Events

Three events $A, B$, and $C$ are pairwise-independent if

$$
\begin{aligned}
& \operatorname{Pr}(A \cap B)=\operatorname{Pr}(A) \cdot \operatorname{Pr}(B) ; \\
& \operatorname{Pr}(B \cap C)=\operatorname{Pr}(B) \cdot \operatorname{Pr}(C) ; \\
& \operatorname{Pr}(A \cap C)=\operatorname{Pr}(A) \cdot \operatorname{Pr}(C) .
\end{aligned}
$$

Three events $A, B$ and $C$ are mutually independent if they are pairwise independent and

$$
\operatorname{Pr}(A \cap B \cap C)=\operatorname{Pr}(A) \cdot \operatorname{Pr}(B) \cdot \operatorname{Pr}(C) .
$$

- Succinct form of the definition of mutual independence:

Three events $A_{1}, A_{2}$, and $A_{3}$ are mutually independent if for all nonempty sets $S \subseteq\{1,2,3\}$,

$$
\operatorname{Pr}\left(\bigcap_{i \in S} A_{i}\right)=\prod_{i \in S} \operatorname{Pr}\left(A_{i}\right) .
$$

## Independence: Examples

1. Experiment: toss a fair coin twice

- $A=$ event that $1^{\text {st }}$ toss is $H$
- $B=$ event that $2^{\text {nd }}$ toss is $H$
- $C=$ event that tosses are different

We showed: $A, B, C$ are pairwise independent, but not mutually independent
2. Experiment: toss a fair coin 3 times

- $A_{i}=$ event that $i$-th toss is $H$ for $i=1,2,3$

Then $\operatorname{Pr}\left(A_{1} \cap A_{2}\right)=\operatorname{Pr}(\{H H T, H H H\})=\frac{1}{4}$

$$
=\frac{1}{2} \cdot \frac{1}{2}=\operatorname{Pr}\left(A_{1}\right) \cdot \operatorname{Pr}\left(A_{2}\right)
$$

Similarly, for $A_{1}$ and $A_{3}$ and for $A_{2}$ and $A_{3}$
$\operatorname{Pr}\left(A_{1} \cap A_{2} \cap A_{3}\right)=\frac{1}{8}=\frac{1}{2} \cdot \frac{1}{2} \cdot \frac{1}{2}=\operatorname{Pr}\left(A_{1}\right) \cdot \operatorname{Pr}\left(A_{2}\right) \cdot \operatorname{Pr}\left(A_{3}\right)$
Conclusion: $A_{1}, A_{2}, A_{3}$ are mutually independent

## Mutually Independent Events

- We can generalize the definition to any number of events.


## Definition: Mutually Independent Events

Events $A_{1}, A_{2}, \ldots, A_{n}$ are mutually independent if for all nonempty sets $S \subseteq\{1,2, \ldots, n\}$,

$$
\operatorname{Pr}\left(\bigcap_{i \in S} A_{i}\right)=\prod_{i \in S} \operatorname{Pr}\left(A_{i}\right) .
$$

Example: toss a fair coin $n$ times

- $A_{i}=$ event that $i$-th toss is $H$ for $i=1,2, \ldots, n$
$A_{1}, A_{2}, \ldots, A_{n}$ are mutually independent
- Independence is a very strong assumption
- Many incorrect arguments in probability are due to
- incorrect assumptions on the independence of events
- not accounting properly for all of the information available
- Need to make sure we are conditioning on the right events


## People v. Collins

## Robbery case in 1968 in a suburb of LA

- A purse was snatched from an elderly person
- A couple seen running from the scene were described as a black man with a beard and a mustache and a blond girl with hair in a ponytail
- Witnesses said they drove off in a partly yellow car

Malcolm and Janet Collins were arrested

- He was black and, though clean shaven when arrested, had evidence of recently having had a beard and a mustache
- She was blond and usually wore her hair in a ponytail
- They drove a partly yellow Lincoln


## https://en.wikipedia.org/wiki/People v. Collins

## People v. Collins

- The prosecution called a mathematics instructor as a witness, who testified about the multiplication rule for Pr
- The prosecutor suggested to the jury to consider the following odds that the defenders were not the robbers

| Black man with beard | 1 in 10 |
| :--- | :--- |
| Man with mustache | 1 in 4 |
| White woman with pony tail | 1 in 10 |
| White woman with blond hair | 1 in 3 |
| Yellow motor car | 1 in 10 |
| Interracial couple in car | 1 in 1,000 |

- What do you think they ruled?
- Do you agree with their decision?


## People v. Collins

- Modeling assumption: each couple matches the witnesses’ description with probability $p$, independently of the other couples
- Prosecutor's argument:
$p=\operatorname{Pr}($ "black man with beard") $\cdot \operatorname{Pr}($ "man with mustache") $\cdot \ldots$

$$
=\frac{1}{10} \cdot \frac{1}{4} \cdot \frac{1}{10} \cdot \frac{1}{3} \cdot \frac{1}{10} \cdot \frac{1}{1000}=\frac{1}{12,000,000}
$$

$\operatorname{Pr}($ "accused couple are not the robbers") $=p$

## 65 237 <br> People v. Collins

The culprit had big eyes, big ears, and big teeth. Statistically, the chance of it being anyone other than the defendant


From http://www.stus.com/stus-cartoon.php?name=People+v.+Collins\&cartoon=evi0007

## People v. Collins: Flaws

- Prosecutor's argument:
$p=\operatorname{Pr}($ "black man with beard") $\cdot \operatorname{Pr}$ ("man with mustache") $\cdot \ldots$

$$
=\frac{1}{10} \cdot \frac{1}{4} \cdot \frac{1}{10} \cdot \frac{1}{3} \cdot \frac{1}{10} \cdot \frac{1}{1000}=\frac{1}{12,000,000}
$$

What should this term be by the Product Rule?

- This argument assumes that events are mutually independent

Is it reasonable to assume, e.g., that "black man with beard" and "man with mustache" are independent events?

## People v. Collins: Flaws

- Prosecutor's argument:
$\operatorname{Pr}$ ("accused couple are not the robbers") $=\operatorname{Pr}($ "a couple matches the description")


## Prosecutor's Fallacy

"The odds of finding this evidence on an innocent man are so small that the jury can safely disregard the possibility that this defendant is innocent"

## People v. Collins: Flaws

- Prosecutor's argument:
$\operatorname{Pr}$ ("accused couple are not the robbers") $=\operatorname{Pr}($ "a couple matches the description")
- This argument fails to condition on all the available information

We need to consider the conditional probability
$\operatorname{Pr}$ ("there are at least 2 couples matching description" |
"there is at least 1 couple matching description")

This information was revealed, so we need to condition on it.

## People v. Collins

- Suppose there are $n$ couples in the LA area, each matching the description with probability $p$, independently of the others
- Let $X=$ number of couples matching the description $\operatorname{Pr}(X \geq 2 \mid X \geq 1)=$


## People v. Collins

- Suppose there are $n$ couples in the LA area, each matching the description with probability $p$, independently of the others
- Let $X=$ number of couples matching the description Example: $n=5,000,000 \quad p=\frac{1}{12,000,000}$

$$
\begin{aligned}
& \operatorname{Pr}(X \geq 1)=1-(1-p)^{n} \approx 0.340759 \\
& \operatorname{Pr}(X \geq 2)=1-(1-p)^{n}-n p(1-p)^{n-1} \approx 0.0660758
\end{aligned}
$$

$$
\operatorname{Pr}(X \geq 2 \mid X \geq 1)=\frac{\operatorname{Pr}(X \geq 2)}{\operatorname{Pr}(X \geq 1)} \approx 0.193907
$$

## People v. Collins

- Suppose there are $n$ couples in the LA area, each matching the description with probability $p$, independently of the others
- Let $X=$ number of couples matching the description

$$
\text { Example: } n=5,000,000 \quad p=\frac{1}{12,000,000}
$$

$$
6.6 \% \text { : not large, but }
$$

$$
\begin{aligned}
& \operatorname{Pr}(X \geq 1)=1-(1-p)^{n} \approx 0.340759 \quad \text { not small either } \\
& \operatorname{Pr}(X \geq 2)=1-(1-p)^{n}-n p(1-p)^{n-1} \approx 0.0660758
\end{aligned}
$$

$$
\operatorname{Pr}(X \geq 2 \mid X \geq 1)=\frac{\operatorname{Pr}(X \geq 2)}{\operatorname{Pr}(X \geq 1)} \approx 0.193907
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## People v. Collins

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$$
\operatorname{Pr}(X \geq 2 \mid X \geq 1)=\frac{\operatorname{Pr}(X \geq 2)}{\operatorname{Pr}(X \geq 1)} \approx 0.193907
$$

19.39\%: that's pretty high

## People v. Collins

- The California Supreme Court overruled the initial guilty verdict
- There are at least two flaws in the prosecution's argument:
- Independence assumptions: "black man with beard" and "man with mustache" may not be independent
- Not conditioning on all the available information ("at least one such couple")

