

Randomness in Computing



LECTURE 5

Last time

- Random variables
- Linearity of expectation
- Jensen's inequality

Today

- Bernoulli and binomial RVs
- Conditional expectation
- Branching process
- Geometric RVs



Bernoulli random variables

- A Bernoulli random variable with parameter p:
 - $\begin{cases} 1 & \text{with probability } p; \\ 0 & \text{otherwise.} \end{cases}$
- The expectation of a Bernoulli R.V. X is $\mathbb{E}[X] = p \cdot 1 + (1 - p) \cdot 0 = p.$

The distribution of a Bernoulli random variable is called the Bernoulli distribution.



Binomial random variables

The binomial distribution with parameters n and p, denoted Bin(n, p), is the distribution of the number of HEADS in n tosses of a coin with bias p. A random variable $X \sim Bin(n, p)$ is a binomial R.V.

Notation ~: ``has probability distribution'' or ``is distributed according to''

• Lemma. The probability distribution of $X \sim Bin(n, p)$ is

$$\Pr[X = j] = \binom{n}{j} p^j (1 - p)^{n-j}$$
 for all $j = 0, 1, ..., n$.

• Lemma. The expectation of $X \sim \text{Bin}(n, p)$ is $\mathbb{E}[X] = np$.

CS 537

Review Question

Throw m balls into n bins uniformly and independently at random.

Let X be the number of balls that land into bin 1.

(Recall that Bin(n,p) is the binomial distribution, i.e., the distribution of the number of HEADS in n tosses of a coin with bias p.)

Then the distribution of X is

- A. Bin (n, m)
- B. Bin (m, 1/n)
- C. Bin $\left(m, \frac{n-1}{n}\right)$
- D. a binomial distribution, but none of the above.
- E. not a binomial distribution.

CS 537

Review Question

Throw *m* balls into *n* bins. Let Y be the number of empty bins.

Compute $\mathbb{E}[Y]$.

- **A.** 1
- B. n/2

C.
$$\left(1-\frac{1}{n}\right)^m$$

$$\mathbf{D.} \quad n\left(1-\frac{1}{n}\right)^m$$

E. None of the above.

Review question

For arbitrary random variables X and Y, by linearity of expectation:

- A. $\mathbb{E}[X + Y] = \mathbb{E}[X] + \mathbb{E}[Y]$
- **B.** $\mathbb{E}[aX + bY] = a\mathbb{E}[X] + b\mathbb{E}[Y]$ for all $a, b \in \mathbb{R}$.
- C. $\mathbb{E}[XY] = \mathbb{E}[X] \mathbb{E}[Y]$
- D. Both A and B are correct.
- E. A,B and C are correct.



Product of independent RVs

• Theorem. For any two **independent** random variables X and Y on the same probability space,

$$\mathbb{E}[X \cdot Y] = \mathbb{E}[X] \cdot \mathbb{E}[Y].$$

• Note. The equality does not hold, in general, for dependent random variables.

Example. We toss two coins.

Let X= number of HEADS, Y= number of TAILS.

Calculate $\mathbb{E}[X]$, $\mathbb{E}[Y]$ and $\mathbb{E}[XY]$.



Conditional expectation: definition

• For any random variable X and event E, the conditional expectation of X given E is

$$\mathbb{E}[X|E] = \sum_{i} i \cdot \Pr[X = i|E],$$

where the sum is over all possible values *i* taken by X.

Exercise

You roll two dice. Let X_1 be the value on the first die, X_2 be the value on the second die, and $X = X_1 + X_2$.

Calculate $\mathbb{E}[X_1 | X = 5]$



Linearity of conditional expectation

• Theorem. For all random variables X and Y and all events A,

$$\mathbb{E}[X + Y \mid A] = \mathbb{E}[X \mid A] + \mathbb{E}[Y \mid A].$$

Also, for all $c \in \mathbb{R}$,

$$\mathbb{E}[cX \mid A] = c \cdot \mathbb{E}[X \mid A].$$

Exercise

You roll two dice. Let A be the event that you got no sixes.

Let X_1 be the value on the first die, X_2 be the value on the second die, and $X = X_1 + X_2$.

Calculate $\mathbb{E}[X \mid A]$



Recall: Law of total probability

Let A be an event and let $E_1, ..., E_n$ be mutually disjoint events whose union is Ω .

Law of total probability.

$$\Pr[A] = \sum_{i \in [n]} \Pr[A \cap E_i] = \sum_{i \in [n]} \Pr[A \mid E_i] \cdot \Pr[E_i].$$



Law of total expectation

Let X be a random variable over sample space Ω and let E_1, \dots, E_n be mutually disjoint events whose union is Ω .

• Law of total expectation:

$$\mathbb{E}[X] = \sum_{i \in [n]} \mathbb{E}[X \mid E_i] \cdot \Pr[E_i].$$

Notable special case of law of total expectation.
 For any two random variables X and Y,

$$\mathbb{E}[X] = \sum_{y \in Range(Y)} \mathbb{E}[X \mid Y = y] \Pr[Y = y].$$



Conditional expectation: definition

For random variables X and Y,

the conditional expectation of X given Y, denoted $\mathbb{E}[X|Y]$,

is a random variable that depends on Y. Its value, when Y = y, is $\mathbb{E}[X | Y = y]$.

• Example: Let *N* be the number you get when you roll a die. You roll a fair coin *N* times and get *H* heads.

Find $\mathbb{E}[H|N]$.

$$\mathbb{E}[H|N=n] = n/2.$$

$$\mathbb{E}[H|N] = N/2.$$



Law of total expectation: compact form

• Recall: For any two random variables X and Y,

$$\mathbb{E}[X] = \sum_{y \in Range(Y)} \mathbb{E}[X \mid Y = y] \Pr[Y = y].$$

• In other words,

$$\mathbb{E}[X] = \mathbb{E}\big[\mathbb{E}[X|Y]\big].$$

• Example: Let *N* be the number you get when you roll a die. You roll a fair coin *N* times and get *H* heads.

Find $\mathbb{E}[H]$.

$$\mathbb{E}[H] = \mathbb{E}\left[\mathbb{E}[H|N]\right] = \mathbb{E}\left[\frac{N}{2}\right] = \frac{3.5}{2} = 1.75.$$



Law of total expectation: application

Branching Process: A program P tosses n coins with bias p and calls itself recursively for every HEADS.

If we call P once, what is total expected number of calls to P that will be generated?

Branching process

Idea: Consider ``generations'' of calls

- Original call is *generation* 0.
- A recursive call is generation i if it was called by a call of generation i-1.

Random variables: $Y_i = \#$ of recursive calls of generation i, for i = 1, 2, ...

Need to find: $\mathbb{E}[Y]$, where $Y = \sum_{i=0}^{\infty} Y_i$.

$$Y_0 = 1$$

$$Y_1 \sim Bin(n, p)$$
, so $\mathbb{E}[Y_1] =$

By compact form of law of total expectation, $\mathbb{E}[Y_2] = \mathbb{E}[\mathbb{E}[Y_2|Y_1]]$

$$\mathbb{E}[Y_2|Y_1=y_1]=\mathbb{E}[Bin(\quad,\quad)]=$$

$$\mathbb{E}[Y_2|Y_1]=$$

$$\mathbb{E}[Y_2] =$$

Similarly,
$$\mathbb{E}[Y_i] = \mathbb{E}[\mathbb{E}[Y_i|Y_{i-1}]] =$$

Similarly,
$$\mathbb{E}[Y_i] = \mathbb{E}[\mathbb{E}[Y_i|Y_{i-1}]] =$$

By linearity of expectation, $\mathbb{E}[Y] = \sum_{i=0}^{\infty} \mathbb{E}[Y_i] = \sum_{i=0}^{\infty} (np)^i = \begin{cases} \frac{1}{1-np} & \text{if } np < 1\\ \text{unbounded } o.w. \end{cases}$



Geometric random variables

The geometric distribution with parameter p, denoted Geom(p), is the distribution of the number of tosses of a coin with bias p until it shows HEADS.

• Lemma. The probability distribution of $X \sim \text{Geom}(p)$ is

$$Pr[X = n] = (1 - p)^{n-1}p$$

for all $n = 1, 2,$

• Lemma. The expectation of $X \sim \text{Geom}(p)$ is $\mathbb{E}[X] = 1/p$.