

# *Randomness in Computing*

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CS  
537

## **LECTURE 16**

**Last time before review**

- Poisson approximation

**Today**

- Discuss midterm
- Review Poisson approximation
- Application: max load
- Application: Coupon Collector
- Random graphs

Let  $X$  and  $Y$  be random variables that take values in  $\mathbb{N}$ .

Which of the following are NOT well defined?

- A.  $\Pr[X]$
- B.  $\Pr[X = Y]$
- C.  $\Pr[X \cap Y]$
- D.  $\Pr[X \text{ is a geometric RV with parameter } p]$
- E.  $\mathbb{E}[\mathbb{E}[X]]$
- F.  $\mathbb{E}[X = 0]$

Make sure your objects are of the right type!

- A **Poisson random variable** with parameter  $\mu$  is given by the following distribution on  $j = 0, 1, 2, \dots$

$$\Pr[X = j] = \frac{e^{-\mu} \mu^j}{j!}$$

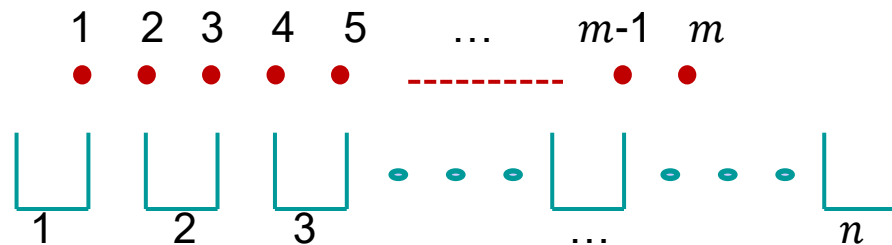
- $\mathbb{E}[X] = \text{var}[X] = \mu$
- Sum of two Poisson RVs is a Poisson RV
- A Poisson RV satisfies Chernoff-type bounds.
- Poisson distribution is a limit of binomial distribution

# Review: Poisson approximation

- The Balls-and-Bins model has dependences.
- The Poisson approximation gets rid of dependencies.

- $m$  balls into  $n$  bins u.i.r.

For  $i \in [n]$ , let



(real world)  $X_i^{(m)} = \#$  of balls in bin  $i$

(Poisson world)  $Y_i^{(m)} \sim \text{Poisson}(\mu)$ , where  $\mu = \frac{m}{n}$  and

$Y_i^{(m)}$  are mutually independent.

- If we condition the Poisson distribution on producing exactly  $k$  balls, then it's the same as the distribution resulting from throwing  $k$  balls into  $n$  bins.

## Poisson Approximation Theorem

Let  $f(x_1, \dots, x_n) \geq 0$  for all  $x_1, \dots, x_n \in \{0, 1, 2, \dots\}$ . Then

$$\mathbb{E} \left[ f \left( X_1^{(m)}, \dots, X_n^{(m)} \right) \right] \leq e\sqrt{m} \cdot \mathbb{E} \left[ f \left( Y_1^{(m)}, \dots, Y_n^{(m)} \right) \right].$$

- **Poisson case:** # of balls in each bin is independent Poisson  $\left(\frac{m}{n}\right)$
- **Corollary.** Any event that has probability  $p$  in the Poisson case has probability  $\leq p \cdot e\sqrt{m}$  in the exact case.

**Proof:** Let  $X$  be the indicator for that event.

Then  $\mathbb{E}[X]$  is the probability that event occurs.

- **Improvements to Theorem and Corollary**

If  $\mathbb{E} \left[ f \left( X_1^{(m)}, \dots, X_n^{(m)} \right) \right]$  is monotonically nonincreasing (or nondecreasing) in  $m$ , then  $e\sqrt{m}$  can be changed to 2.

## *n balls into n bins*

- Before (by Chernoff):  $\Pr \left[ \text{MaxLoad} > \frac{3 \ln n}{\ln \ln n} \right] \leq \frac{1}{n}$  for s.l.  $n$
- Theorem.  $\Pr \left[ \text{MaxLoad} < \frac{\ln n}{\ln \ln n} \right] \leq \frac{1}{n}$  for s.l.  $n$

Proof: Let  $M = \frac{\ln n}{\ln \ln n}$

# Application: Coupon Collector

$X = \#$  of coupons observed before obtaining 1 of each of  $n$  types

- Before:  $\mathbb{E}[X] =$  and  $\Pr[X > n \ln n + cn] \leq e^{-c} \forall c > 0$

Review:  $\Pr$ [not obtaining coupon  $i$  in  $n \ln n + cn$  steps] is

- Theorem.  $\Pr[X > n \ln n + cn] \leq 2(1 - e^{-1.5 \cdot e^{-c}}) \forall c > 0$ , s.l.  $n$
- MU:  $\lim_{n \rightarrow \infty} \Pr[X > n \ln n + cn] = 1 - e^{-e^{-c}} \forall c > 0$

# Application: Coupon Collector

*$X = \#$  of coupons observed before obtaining 1 of each of  $n$  types*

- Theorem.  $\Pr[X > n \ln n + cn] \leq 2(1 - e^{-1.5 \cdot e^{-c}}) \quad \forall c > 0, \text{ s.t. } n$

Proof: Balls and bins view

$n$  bins

- $X = \#$  balls thrown before all bins nonempty
- Consider  $m = n(\ln n + c)$  balls.
- Let  $B =$  event that there is an empty bin.

$$\Pr[X > n \ln n + cn] = \Pr[B]$$

- Idea: Use Poisson approximation:

# balls in each bin is Poisson( $\mu$ ) with  $\mu = \ln n + c$

- $E_i =$  event that bin  $i$  is empty. Then

- $\Pr[E_i] = e^{-\mu} = e^{-(\ln n + c)} = \frac{e^{-c}}{n}$



# Application: Coupon Collector

**Theorem.**  $\Pr[X > n \ln n + cn] \leq 2(1 - e^{-1.5 \cdot e^{-c}}) \quad \forall c > 0, \text{ s.t. } n$

**Proof:**  $n$  bins and  $m = n(\ln n + c)$  balls.

- Let  $B =$  event that there is an empty bin:  $\Pr[X > n \ln n + cn] = \Pr[B]$
- Poisson approximation: # balls in each bin is  $\text{Poisson}(\mu)$ ,  $\mu = \ln n + c$
- $E_i =$  event that bin  $i$  is empty. Then  $\Pr[E_i] = \frac{e^{-c}}{n}$

$$\begin{array}{c}
 \begin{array}{|c|} \hline \mathbf{0 \leq} \\ \hline \end{array}
 \begin{array}{|c|} \hline \mathbf{for} \\ \hline \end{array}
 \begin{array}{|c|} \hline \mathbf{x \leq 1/2} \\ \hline \end{array} \\
 \swarrow \quad \searrow \\
 \begin{array}{|c|} \hline \mathbf{e^{-1.5x} = e^{-x-0.5x} \leq} \\ \hline \end{array}
 \begin{array}{|c|} \hline \mathbf{e^{-x-x^2} \leq 1-x \leq e^{-x}} \\ \hline \end{array}
 \end{array}$$

$G_{n,p}$  :  $n$  vertices, each of  $\binom{n}{2}$  potential edges is present w.p.  $p$

$G_{n,M}$  :  $n$  vertices, a uniformly random  $M$ -edge graph

Suppose  $G \sim G_{n,p}$

What is the expected number of edges in  $G$ ?

- A.  $p$
- B.  $n/p$
- C.  $pn$
- D.  $pn^2$
- E.  $pn(n-1)/2$
- F. None of the above

The distribution  $G_{n,p}$  conditioned on producing a graph with  $M$  edges is

- A.  $G_{n,M}$
- B.  $G_{n,pn}$
- C.  $G_{n,p} \binom{n}{2}$
- D. None of the above

$G_{n,p}$  :  $n$  vertices, each of  $\binom{n}{2}$  potential edges is present w.p.  $p$

$G_{n,M}$  :  $n$  vertices, a uniformly random  $M$ -edge graph

- When  $p = \frac{M}{\binom{n}{2}}$ , the number of edges in a graph  $\sim G_{n,p}$  is concentrated around  $M$ .
- **Conditioning claim.**  $G_{n,p}$  conditioned on having  $M$  edges  $\Leftrightarrow G_{n,M}$

**Graph property:** whether a graph has it doesn't depend on how its vertices are labeled (it holds for  $G \Leftrightarrow$  it holds for all graphs isomorphic to  $G$ ).

A graph property is **monotone increasing** if:

for every graph  $G$  that has the property, every graph  $G'$  obtained by *adding* edge(s) to  $G$  also has the property.

**Examples:** connected, has a component with  $\geq k$  vertices, has a cycle, ...

A graph property is **monotone decreasing** if:

for every graph  $G$  that has the property, every graph  $G'$  obtained by *deleting* edge(s) from  $G$  also has the property.

**Examples:** acyclic, bipartite, is a forest...

**Not monotone:** e.g., is a tree.

## Theorem

Consider a monotone increasing graph property.

Let  $P(n, M)$  be the probability it holds for a graph  $\sim G_{n,M}$ .

Let  $\tilde{P}(n, p)$  be the probability it holds for a graph  $\sim G_{n,p}$ .

Fix a constant  $\epsilon \in (0,1)$ . Let  $p^- = (1 - \epsilon) \cdot \frac{M}{\binom{n}{2}}$  and  $p^+ = (1 + \epsilon) \cdot \frac{M}{\binom{n}{2}}$ . Then

$$\tilde{P}(n, p^-) - e^{-\Omega(M)} \leq P(n, M) \leq \tilde{P}(n, p^+) + e^{-\Omega(M)}$$

**Proof:** Let R.V.  $X = \#$  of edges in a graph chosen  $\sim G_{n,p^-}$ .

$$\tilde{P}(n, p^-) = \sum_{k=0}^{\binom{n}{2}} P(n, k) \cdot \Pr[X = k]$$

**By Law of Total Probability  
+ Conditioning Claim**

$$= \sum_{k \leq M} P(n, k) \cdot \Pr[X = k] + \sum_{k > M} P(n, k) \cdot \Pr[X = k]$$

$$\leq P(n, M) \cdot \Pr[X \leq M] + \Pr[X > M]$$

Mon. increasing property  $\Rightarrow$   
 $P(n, k) \leq P(n, M) \forall k \leq M$

$$\leq P(n, M) + \Pr[X > M]$$

## Theorem

Consider a monotone increasing graph property.

Let  $P(n, M)$  be the probability it holds for a graph  $\sim G_{n,M}$ .

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**Proof:** Let R.V.  $X = \#$  of edges in a graph chosen  $\sim G_{n,p^-}$ .

$$\tilde{P}(n, p^-) \leq P(n, M) + \Pr[X > M]$$

Apply Chernoff:

$$\mu = \mathbb{E}[X] = (1 - \epsilon)M \stackrel{\geq 1 + \epsilon}{\text{since } \epsilon \in (0, 1)}$$

$$\begin{aligned} \Pr[X > M] &= \Pr\left[X > \frac{1}{1 - \epsilon} \cdot \mu\right] \leq \Pr[X > (1 + \epsilon) \cdot \mu] \\ &\leq e^{-\frac{\epsilon^2 \mu}{3}} = e^{-\frac{\epsilon^2 (1 - \epsilon) M}{3}}. \end{aligned}$$



## Theorem

Consider a monotone increasing graph property.

Let  $P(n, M)$  be the probability it holds for a graph  $\sim G_{n,M}$ .

Let  $\tilde{P}(n, p)$  be the probability it holds for a graph  $\sim G_{n,p}$ .

Fix a constant  $\epsilon \in (0,1)$ . Let  $p^- = (1 - \epsilon) \cdot \frac{M}{\binom{n}{2}}$  and  $p^+ = (1 + \epsilon) \cdot \frac{M}{\binom{n}{2}}$ . Then

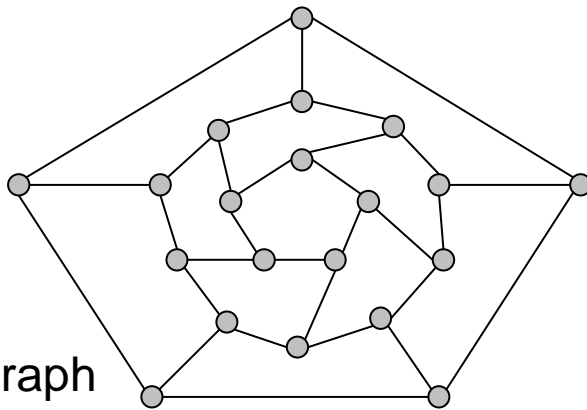
$$\tilde{P}(n, p^-) - e^{-\Omega(M)} \leq P(n, M) \leq \tilde{P}(n, p^+) + e^{-\Omega(M)}$$

**Proof:** The proof for the second inequality is similar.

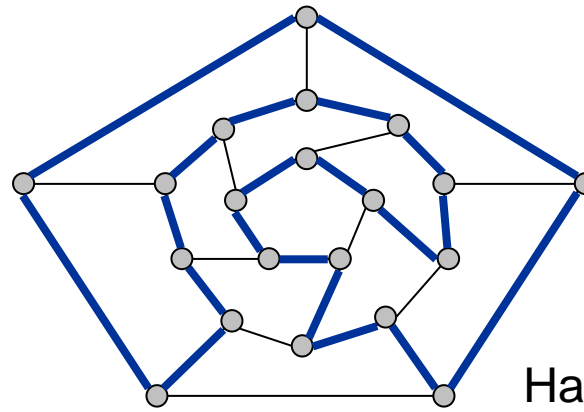
A similar theorem holds for monotone decreasing properties.

# Finding Hamiltonian cycles

A **Hamiltonian cycle** is a cycle that visits each vertex exactly once.



Input graph



Hamiltonian cycle

- Finding Hamiltonian cycles is NP-hard.

The property of having a Hamiltonian cycle is

- A. Monotone increasing
- B. Monotone decreasing
- C. Not monotone

# CS 537 | Finding Hamiltonian cycles in random graphs

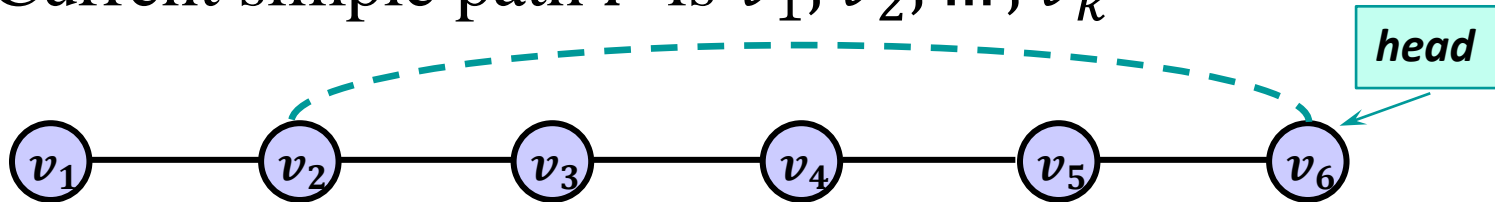
## Main Theorem

Suppose  $p \geq \frac{40 \ln n}{n}$ . There is a polynomial time randomized algorithm that, given  $p$  and  $G \sim G_{n,p}$ , finds a Hamiltonian cycle in  $G$  with probability  $1 - O\left(\frac{1}{n}\right)$ .

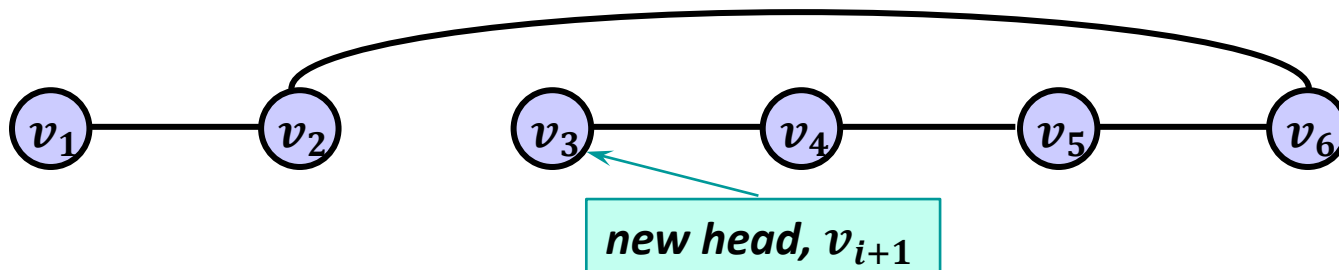
**Corollary:** Hamiltonian cycle exists in  $G \sim G_{n,p}$  w. p.  $1 - O\left(\frac{1}{n}\right)$ .

# Rotation operation

- Current simple path  $P$  is  $v_1, v_2, \dots, v_k$



- **Rotation** of  $P$  with an edge  $(v_k, v_i)$  in  $G$ ,  $i \in [k - 1]$



- If  $i = k - 1$ , no change.
- If  $k = n$  and  $i = 1$ ,  
rotation edge  $(v_k, v_i)$  **closes** Hamiltonian path.